1 **Title:** Challenges of COVID-19 Case Forecasting in the US, 2020-2021

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65 Abstract

66 During the COVID-19 pandemic, forecasting COVID-19 trends to support planning and response was a priority for 67 scientists and decision makers alike. In the United States, COVID-19 forecasting was coordinated by a large 68 group of universities, companies, and government entities led by the Centers for Disease Control and Prevention 69 and the US COVID-19 Forecast Hub (https://covid19forecasthub.org). We evaluated approximately 9.7 million 70 forecasts of weekly state-level COVID-19 cases for predictions 1-4 weeks into the future submitted by 24 teams 71 from August 2020 to December 2021. We assessed coverage of central prediction intervals and weighted 72 interval scores (WIS), adjusting for missing forecasts relative to a baseline forecast, and used a Gaussian 73 generalized estimating equation (GEE) model to evaluate differences in skill across epidemic phases that were 74 defined by the effective reproduction number. Overall, we found high variation in skill across individual models, 75 with ensemble-based forecasts outperforming other approaches. Forecast skill relative to the baseline was 76 generally higher for larger jurisdictions (e.g., states compared to counties). Over time, forecasts generally 77 performed worst in periods of rapid changes in reported cases (either in increasing or decreasing epidemic

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78 phases) with 95% prediction interval coverage dropping below 50% during the growth phases of the winter 79 2020, Delta, and Omicron waves. Ideally, case forecasts could serve as a leading indicator of changes in 80 transmission dynamics. However, while most COVID-19 case forecasts outperformed a naïve baseline model, 81 even the most accurate case forecasts were unreliable in key phases. Further research could improve forecasts 82 of leading indicators, like COVID-19 cases, by leveraging additional real-time data, addressing performance 83 across phases, improving the characterization of forecast confidence, and ensuring that forecasts were coherent 84 across spatial scales. In the meantime, it is critical for forecast users to appreciate current limitations and use a 85 broad set of indicators to inform pandemic-related decision making.

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87 Author Summary

88 As SARS-CoV-2 began to spread throughout the world in early 2020, modelers played a critical role in predicting 89 how the epidemic could take shape. Short-term forecasts of epidemic outcomes (for example, infections, cases, 90 hospitalizations, or deaths) provided useful information to support pandemic planning, resource allocation, and 91 intervention. Yet, infectious disease forecasting is still a nascent science, and the reliability of different types of 92 forecasts is unclear. We retrospectively evaluated COVID-19 case forecasts, which were often unreliable. For 93 example, forecasts did not anticipate the speed of increase in cases in early winter 2020. This analysis provides 94 insights on specific problems that could be addressed in future research to improve forecasts and their use. 95 Identifying the strengths and weaknesses of forecasts is critical to improving forecasting for current and future 96 public health responses.

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98 Introduction

99 Predicting the trajectory of an epidemic to support control and mitigation planning is the primary objective of
 100 infectious disease forecasting. To this end, large-scale, collaborative forecasting efforts across multiple disease

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101 systems, such as influenza (1-3), dengue (4), West Nile (5), and Ebola viruses (6), have been integrated into 102 routine public health workflows and emergency response (7). Researchers in academia, private institutions, and 103 the United States (US) government built upon these frameworks to incorporate forecasts into the COVID-19 104 information systems used to inform pandemic response and created the US COVID-19 Forecast Hub. In April 105 2020, the US Centers for Disease Control and Prevention (CDC) and the COVID-19 Forecast Hub began collecting 106 COVID-19 death forecasts (8). Compared to death reports, case reports are a leading indicator of SARS-CoV-2 107 infections, as the time from infection to case report is typically shorter than that between infection and death 108 report. Hence, information gleaned from case forecasts is potentially more actionable.

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110 Case forecasts for all US counties (n=3,143), states (n=50), territories (n=5), the District of Columbia (DC), and 111 the nation as a whole were generated and collected beginning in July 2020, with ensemble forecasts of cases 112 first posted on a CDC webpage on August 6, 2020 (8,9). Because of their potential utility, case forecasts were 113 also integrated into US government web pages and situational awareness updates (10). In addition, county-level 114 case forecasts were used to inform vaccine trial site selection (11) and COVID-19 case forecasts have been cited 115 as useful for guiding personal risk-based decisions (12). Because these forecasts influence policies and personal 116 decisions, accuracy and precision of the forecasts is of the utmost importance. Incorrect forecasts can lead to 117 inappropriate policy implementation and resource allocation, and also to erosion of trust in public health 118 institutions (13).

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As part of routine use of the case forecasts in the COVID-19 response, real-time evaluation was conducted. One of the performance metrics included in the evaluation was the 95% prediction interval (PI) coverage, an estimate of the frequency at which the interval captures the eventually observed data. The 95% PI of a reliable forecast should capture eventually reported cases 95% of the time. However, the real-time evaluation indicated that case forecasts were not always reliable, with much lower 95% PI coverage than expected (14). For example, in November 2020 as the 2020-2021 winter wave began, the 95% PI coverage for all states and territories was less

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than 50% for even the shortest, 1-week ahead forecasts from the ensemble – generally the most reliable forecast. Repeated periods of low coverage during subsequent surges led CDC to stop posting COVID-19²Case forecasts in December 2021. Though these forecasts showed poor performance, there are opportunities to develop more precise and reliable future predictions.

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131 Evaluation of forecast performance provides an opportunity not only to assess prediction skill for the purposes 132 of improving forecasts, but also to assess the reliability of the forecasts and foster transparency between 133 forecast users and creators. While evaluation is recommended in forecasting research guidelines (i.e., EPIFORGE 134 2020 (15), a systematic review of COVID-19 models showed that half of published models did not include 135 probabilistic predictions and that approximately one-fourth of published models did not include performance 136 evaluations (16). We have previously evaluated forecast performance of cumulative (17) and incident (18) 137 COVID-19 deaths submitted to the COVID-19 Forecast Hub. Given that an ensemble of submitted models 138 provided consistently accurate probabilistic forecasts at different scales in both evaluations, here we apply 139 similar methods to assess the prediction skill of the COVID-19 case forecasters, with particular interest in the 140 COVIDhub ensemble model (that is, a model that combine predictions from forecasts submitted to the Forecast 141 Hub). Specifically, we analyze prediction interval coverage and other aspects of nearly 10 million individual 142 forecasts collected by the COVID-19 Forecast Hub for US jurisdictions between July 2020 and December 2021, 143 the full period over which COVID-19 case forecasts were published by the CDC. We analyze relative forecast 144 performance across spatial scales and phases of the pandemic to identify limitations and opportunities for 145 future improvement of case forecasts.

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147 **Results**

148 Summary of Included Team Forecasts

A total of 14,960,171 forecasts were submitted by 67 teams throughout the analysis period (see Supporting Information [S] 1 for submission patterns over time). Because forecasts were submitted at multiple geographic scales, we stratified analyses for 1) national forecasts, 2) state (including all 50 states), territory (US Virgin Islands and Puerto Rico), and DC forecasts), 3) county level forecasts (include all 3,143 counties and county equivalents), split into five equal sized groups based on county population size.

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155 We first evaluated forecasts for inclusion criteria based on numbers of locations, horizons, and time periods 156 forecast with the same model. Briefly, teams were included if they submitted the full range of required 157 quantiles, included at least 50 of states/territories/DC or 75% of counties, and produced forecasts at least four 158 weeks into the future for at least 50% of the time points in the study period. At the national level, 22 sets of 159 team forecasts met these criteria (5,136 forecasts across dates and forecast horizons), 23 sets of team forecasts 160 met the state/territory level criteria (280,132 forecasts across jurisdictions, dates, and forecast horizons), and 15 161 sets of team forecasts met the county-level criteria (9,415,460 forecasts across counties, dates, and forecast 162 horizons). Overall, 64.8% of all submitted forecasts were included in the analysis (9,700,728 forecasts). Of the 163 included forecasts, 11 sets of team forecasts met the inclusion criteria for analyzing submissions across all 164 geospatial scales (8,125,220 forecasts for specific locations, date and forecast horizon).

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Each team included in the analysis submitted forecasts that were generated from unique model structures, data inputs, and assumptions (S1). Two naïve models (the COVIDhub-baseline and CEID-Walk) and four ensemble models (the COVIDhub-4_week_ensemble, the COVIDhub-trained_ensemble, LNQ-ens1, and UVA-Ensemble), which combined multiple forecasts into one, were included in the 26 models evaluated (see S1 Table 1.1). The COVIDhub-baseline model projects the number of reported cases in the most recent week as the median predicted value for the next 4 weeks. CEID-Walk is a random walk model with a simple method for removing

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outliers. A total of seven models included data on COVID-19 hospitalizations, 12 models incorporated
 demographic data, and seven models used mobility data. Of the 26 evaluated models, three assumed that social
 distancing and other behavioral patterns changed during the prediction period.

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176 The evaluation period consisted of 1-4 week ahead forecasts submitted in the 73 weeks from July 28, 2020 177 through December 21, 2021. Multiple phases of the US epidemic were included: the late summer 2020 increase 178 in several locations, a large late-fall/early-winter surge in 2020/2021, the rise and fall of the Delta variant in the 179 summer and fall of 2021, and the early phase of the Omicron variant's dominance in winter 2021 (Figure 1A). 180 Performance of the national ensemble forecasts varied over this period (Figure 1B). For some forecasts, the 181 median predictions were close to the cases eventually reported, and most reported numbers fell within the 95% 182 Pls. However, forecasts made at other times, such as January 2021 or December 2021, diverged widely from the 183 reported data. At those times, the forecasts missed substantial decreases and increases, respectively, with 184 reported cases falling within the 95% prediction interval for only 1-week ahead forecasts.

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Figure 1. Weekly incident reported COVID-19 cases per 100K population, nationally (in black) and per state/territory/DC (in gray), over time in panel A. Panel B shows a subset of COVIDhub-4_week_ensemble forecasts (in green) over time, with the median predictions represented as lines and points and the 95% prediction intervals in bands. Reported incident cases (counts per week) are shown in gray. In both plots, the black, dashed vertical line shows the date that public communication of the case forecasts was paused.

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192 Aggregate performance

We evaluated aggregate forecast performance with two metrics: Weighted Interval Score (WIS), a proper score considering both precision and accuracy, and prediction interval coverage, an indicator of forecast uncertainty. Lower WIS values reflect forecasts with probability mass closer to observed values. We assessed scaled pairwise

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196	WIS relative to the baseline model (referred to throughout as relative WIS, or rWIS) for national and
197	state/territory/DC forecasts (Figure 2). A rWIS less than one indicates performance that is better than the
198	baseline model.

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Figure 2: Percent of weeks with complete submissions for all sets of team forecasts, scaled, pairwise relative Weighted Interval Score (rWIS; see *Methods* for description), observed 95% prediction interval coverage, by geographical scale of submitted forecasts. Teams are sorted by increasing state/territory/DC rWIS values.

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Overall, seven of 22 team's forecast models outperformed the COVIDhub-baseline model at the state/territory/DC level (i.e., had rWIS values less than 1.0), and 11 outperformed the baseline model at the national level. Six of these teams outperformed the baseline model at both scales: LNQ-ens1, COVIDhub-4 week ensemble, USC-SI kJalpha, LANL-GrowthRate, Microsoft-DeepSTIA, and CU-select.

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PI coverage at the 95% level should be close to 95% for well calibrated forecasts. However, it was lower for most sets of team forecasts, with only one (LNQ-ens1) having coverage of at least 90% at all scales, while others were as low as 23%. PI coverage at 50% and 80% levels were also well below nominal levels for most sets of team forecasts, including the COVIDhub-4_week_ensemble (Figure 3). For the 50% prediction interval, no sets of team forecasts had coverage better than 36% at any scale. Only two sets of team forecasts had better coverage than the COVIDhub-4_week_ensemble for the geographic scales in which they submitted forecasts: LNQ-ens1 (all scales) and JHU_UNC_GAS-StatMechPool (state/territory/DC and large county levels).

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Figure 3: Expected and observed coverage rates for central 50%, 80% and 95% prediction intervals aggregated over time and horizon for national forecasts (panel A), state/territory/DC forecasts (panel B), the largest county forecasts (panel C). The dashed line represents optimal expected-coverage. Team forecasts that had

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closer to nominal coverage than the COVIDhub-4_week_ensemble model at all three coverage levels are
labeled on the right side of the plots.

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Forecast skill also showed distinct patterns across jurisdictional scales, with rWIS decreasing for larger jurisdiction scales (e.g., national vs. state/territory) or population sizes (e.g., larger counties vs. smaller counties, Figure 4) for most sets of team forecasts. In contrast to this general trend, for three sets of team forecasts, that pattern was inverted, one team had no distinct pattern, and the COVIDhub-4_week_ensemble had markedly consistent rWIS across all scales. Consistent with the aggregate findings, both LNQ-ens1 and COVIDhub-4_week_ensemble had rWIS lower than 1.0 at all scales, while LANL-GrowthRate had rWIS greater than 1.0 for smaller counties.

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Figure 4: Scaled, pairwise relative Weighted Interval Score (rWIS) (see *Methods* for description) by spatial scale for sets of team forecasts that submitted forecasts for the US nation, states/territories/DC, and all US counties. WIS is averaged across all horizons. The COVIDhub-baseline model has, by definition, a rWIS of 1 (horizontal dashed line). Teams are ordered by increasing state/territory/DC rWIS with the most accurate model on the left. Points for each team are staggered horizontally to show overlapping WIS values.

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237 **Performance across jurisdictions**

There was additional variability in forecast skill between jurisdictions. Only two team forecasts (LNQ-ens1 and COVIDhub-4_week_ensemble) performed as well as or better than the baseline for all included states and territories (Figure 5). Variation was higher between team forecasts than between specific jurisdictions, but the baseline model tended to outperform more models in some jurisdictions (e.g., the baseline was better in Colorado, Kansas, Puerto Rico) than in others (e.g., the baseline was worse in Mississippi, South Carolina, West Virginia).

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Figure 5: Scaled, pairwise relative Weighted Interval Score (rWIS; see *Methods* for description) by location for national and state/territory/DC forecasts, averaged across all horizons through the entire analysis period. National estimates are displayed first, followed by jurisdictions in alphabetical order. Team forecasts are ordered by increasing average state/territory/DC rWIS.

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250 **Performance over time**

251 While rWIS varied between team forecasts and jurisdictions, it varied even more over time (Figure 6). For 252 example, all models had relatively high WIS in December 2020-January 2021 and low WIS in June 2021. 253 Prediction interval coverage also varied between teams and over time, with most team forecasts exhibiting 254 times of low coverage. Across most time points, the baseline model outperformed many team forecasts, 255 including the COVIDhub-4 week ensemble, though the ensemble more often outperformed the baseline in 256 both metrics at the national, state/territory, and large county scales. Increased WIS and decreased prediction 257 interval coverage generally occurred with increasing case counts, such as in the fall of 2020 and summer of 258 2021. The worst performance was in the early Omicron wave in the winter of 2021. For the last set of ensemble 259 forecasts CDC 2021 (https://www.cdc.gov/coronavirus/2019by in December posted 260 ncov/science/forecasting/forecasts-cases.html), the WIS reached the highest level ever for all scales and the 261 reported case numbers were outside the 95% prediction interval for most locations at every forecast horizon.

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Figure 6: Forecast accuracy over time, aggregated by geographic units, forecast horizon, and prediction date. Panels A-C show average Weighted Interval Score (WIS); panels D-F show 95% prediction interval coverage. The black, dashed vertical line in all panels shows the date that public communication of the case forecasts was paused. The black, dashed horizontal line in panels D-F shows nominal 95% interval coverage. National

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level forecasts are presented in A and D, state/territory/DC forecasts in B and E and large county level
 forecasts in C and F.

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270 To further investigate these temporal patterns in performance, we first classified each forecast week as 271 increasing, peak, decreasing, or nadir based on the estimated time-varying reproduction number for that given 272 week and jurisdiction. We then fitted Gaussian generalized estimating equations (GEE) models for each set of 273 team forecasts, using a normalized, log transformed WIS value per forecast time and location as the model 274 outcome. The regression models were adjusted for each prediction horizon and included a natural spline with 275 two degrees of freedom for the time/state reported case counts to adjust for intrinsic increases in WIS due to 276 higher values in reported cases (see S6). In agreement with the aggregated results (Figure 2), we found that the 277 expected WIS at the mean number of case counts across all jurisdictions was lower than the baseline for the 278 better performing models (6 team forecasts and the ensemble) and higher than the baseline for others (8 team 279 forecasts).

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Forecasts skill also varied across epidemic phases (Figure 7B). Compared to the baseline model across all phases, overall skill for most models was better in nadir and peak phases and worse in increasing and decreasing phases. LNQ-ens1 and the COVIDhub ensemble outperformed the baseline model in all epidemic phases between August 1, 2020 and January 15, 2022, while several other team models outperformed the baseline in some phases.

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Figure 7. Estimated marginal mean Weighted Interval Score (WIS) and 95% confidence intervals for mean cases from team-specific GEE models for all 51 jurisdictions (Panel A). The 95% confidence intervals for the COVIDhub-baseline model are shown in dashed red vertical lines. Panel B presents each team's estimated marginal mean WIS per phase, scaled to the COVIDhub-baseline model's estimated marginal mean WIS for all epidemic phases. Teams with higher estimated marginal mean WIS values (i.e., greater than 1.0) are

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292 presented in shades of orange while teams with lower estimated marginal mean WIS (i.e., less than 1.0) are 293 shown in shades of green. Forecasts for a team in a particular phase are marked with an asterisk (*) if the 80% 294 confidence interval of the expected WIS outcome (normalized and on the log scale) was estimated by a model 295 to be lower than the expected WIS of the COVIDhub-baseline model for all phases.

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297 To examine whether our results were affected by reporting anomalies, we also conducted sensitivity analyses 298 for data revisions, when data were revised at a later date, and for outlier data points, when reported cases were 299 outside of weekly expected ranges (see S2). We first identified weeks in which revised case counts as of April 2, 300 2022 differed from the case counts initially reported for that week, excluded them from the dataset, and reran 301 the GEE models. With this partial dataset, the results were essentially unchanged. Next, we identified outliers as 302 reported case counts outside of the expected range by at least two of the three following algorithms: a rolling 303 median, a seasonal trend decomposition, and a seasonal trend decomposition without a seasonality term, each 304 method over a 21-day window. Approximately 3% of weeks (686 of 27,489 total week-location combinations in 305 the analysis period) had at least one day of reported cases identified as an outlier. We then excluded the weeks 306 with outliers and the week following an outlier and reran the GEE models. This sensitivity analysis had 307 comparable results to the models with the full data (see S2 Figure 2.3, Panel A.).

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309 **Discussion**

We evaluated performance of 9.7 million COVID-19 case forecasts at multiple geospatial scales in the US over approximately a year and a half. Real-time analyses and those presented here revealed important limitations in these forecasts. Forecast prediction intervals were largely over-confident, that is, prediction interval coverage was lower than the nominal value, particularly when case numbers were changing rapidly and forecasts could have been most useful. Few team forecasts outperformed a relatively simple and minimally informative baseline model. Forecast skill degraded for smaller geographic scales where forecasts could potentially be most useful.

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Forecast skill was also lowest when case counts were changing the most, in phases of increasing or decreasing transmission. These limitations of case forecasts indicate key areas for improvement and important reasons to use case forecasts with caution.

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320 Several technical challenges for forecasts were evident in these analyses. First, cases are a relatively early 321 indicator of transmission, with no clear leading signal in traditional public health surveillance data (e.g., unlike 322 for death forecasts, where case counts themselves can provide information for predicting future deaths). While 323 non-traditional data sources may provide a useful predecessor to changing population case counts, the evidence 324 from previous work is unclear. For example, internet searches, medical claims, and online surveys have been 325 used to modestly improve case forecast accuracy relative to models without those data (19). Estimating case 326 counts using both wastewater and clinical surveillance data has shown mixed results (20-23). Additional 327 integration of temporal dynamics could also be helpful. The case forecasts analyzed here were developed and 328 evaluated based on the date when cases were reported, not when individuals were infected, became ill, sought 329 care, or were tested. Additional detail on those dates could enable models to better capture the current 330 dynamics using nowcasting approaches giving earlier signals of change.

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332 Second, and likely related to the challenge of cases being an early indicator, the models had substantial variation 333 in skill between epidemic phases. In general, forecast skill was worst for the increasing phase followed by the 334 decreasing phase. In many of these periods of low performance (e.g., the 2020-2021 winter, Delta, and Omicron 335 waves), the COVIDhub ensemble predicted possible or probable increases or decreases, but not at the rate that 336 actually occurred. This effect may be even stronger than our results show as they rely on a comparison to the 337 baseline which, by definition, does not predict change. While epidemic phase is unknown in real time, it too can 338 be estimated, and these results and others suggest that accounting for epidemic phase when making predictions 339 could improve the forecast skill of ensemble models (24,25). Additional data, as discussed above, or model 340 components associated with distinct phases could also help improve predictive capabilities. Seasonal changes in

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transmission biology and human behavior, emergence of variants, and changing mitigation behavior all contribute to transmission dynamics. While some forecasting models incorporate seasonality and variants, integration of human behavior to characterize the link between behavior and transmission has lagged (13,26– 28). Ensemble approaches offer another opportunity to mitigate phase-specific differences. Team modeling skill across phases was highly heterogeneous, but two ensemble approaches were better than the baseline in all phases.

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348 Another challenge across most forecasts was overconfidence, a pattern seen with other infectious disease 349 forecasts (4,18). The baseline model predicted a flat trend, yet it outperformed many sets of team forecasts in 350 the increasing and decreasing phase only because its predictions had high uncertainty around that flat trend. 351 The COVIDhub ensemble performance, on the other hand, benefitted by combining uncertainty across multiple 352 models, yet, like the constituent models, also exhibited overconfidence. The temporal and phase-specific 353 analyses suggest that it is, during rapid increases and decreases, that model overconfidence is most pronounced. 354 Previous infectious disease forecasting work has shown that ensembles tend to have wider prediction intervals 355 that are more likely to capture the eventually reported outcome and thus reduce overconfidence compared to 356 their constituent models (4,18). Wider prediction intervals, reflecting increased uncertainty, can mediate some 357 impacts of overconfidence. However, forecasts would be most useful if they were both reliable and informative -358 that is, if they could accurately capture the uncertainty, while also providing more precise estimates, rather than 359 merely increased uncertainty (29,30).

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Finally, while forecasts would be most actionable at local scales, performance was generally worse for smaller than larger jurisdictions. Other infectious disease forecasting systems have found better forecast skill at smaller geographic scales, likely because local transmission dynamics (e.g., a county) are a better predictor of local than aggregate transmission (e.g., a state) (31). We compared WIS across scales by comparison to the baseline model to adjust for missing forecasts and for WIS scaling relative to the magnitude of observed outcomes. After those

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adjustments, population size had a clear association with forecast, likely reflecting the relative role of stochastic dynamics. For better local forecasts, models may need to explicitly account for stochasticity. Forecasts could also be improved by better leveraging spatial information, such as dynamics in neighboring counties or nearest urban centers. Local forecasts remain a key public health need, as local forecasts are more likely to reflect local conditions and motivate local mitigation action.

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372 Overall, these findings, as well as the real-time evaluations, indicated that COVID-19 case forecasts were not 373 reliable as a single indicator for pandemic response of a novel pathogen. Similar to other forecasting studies, we 374 found that the ensemble was among the most reliable forecasts (3,4,18,32), outperformed only by LNQ-ens1 375 across the metrics evaluated here. Thus, while the overall best forecasts had poor performance at key times, 376 other forecasts were often even worse at these same time points. Weighted (or trained) ensembles offer 377 another potential avenue for improvement (33-35), but the version implemented here did not outperform the 378 simple, median ensemble, likely reflecting limited historical data (36) and variation in team forecast submissions 379 (37,38).

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381 While COVID-19 deaths are a more lagging indicator of infections than case reports, and so may be less useful as 382 an input to public health decision making, forecasts of deaths have generally been more reliable (18). Similarly, 383 COVID-19 hospitalization forecasts in France have also shown high forecast skill (39). Better performing US death 384 and French hospitalization forecasts share one factor in common: models generally used local case reports as an 385 input to inform their forecasts. While public health decision making should not rely on case forecasts alone, they 386 may still be helpful in the context of other important indicators, such as the case, hospitalizations, and death 387 reports. Nowcasts of reports and real-time estimates of the effective reproductive number can also provide 388 insight on current dynamics (40–43). Together, a suite of indicators is more informative for outbreak response 389 than a leading indicator alone.

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391 The analysis presented here includes important findings about real-time applied forecasting in an emerging 392 pandemic to inform pandemic response rather than to address specific research aims of improving predictions. 393 Several factors limit the strength of our findings and ability to understand underlying mechanisms of predictive 394 performance. Notably, we compared the forecasts to a changing record of reported cases. Throughout the 395 COVID-19 outbreak, cases have been reported with jurisdiction- and time-varying delays and have been revised 396 over time, resulting in varying forecast targets. In addition, the definition of a reported COVID-19 case also 397 changed over time and varied between states. These changes were a result of many factors, including laboratory 398 capacity and implementation of home-based testing, and may have affected forecast skill in other ways. Our 399 sensitivity analyses found no qualitative differences in our main findings when we excluded forecasts for time 400 points with revised data or when we excluded outlier data points. Nevertheless, forecasting teams were greatly 401 impacted by the evolving landscape of COVID-19 case surveillance. More timely and consistent reports likely 402 would improve both the process of making forecasts and forecast skill.

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404 The overall goal of the COVID-19 Forecast Hub was to provide forecasts in near real-time for decision making. 405 While the collaborative efforts of the Hub achieved this goal despite a changing epidemic landscape, 406 nevertheless, the e open nature of COVID-19 forecasting also limits understanding the drivers of forecast 407 performance. Many teams participated at different times, some intermittently, and provided varied and limited 408 descriptions of their forecast methods. While we were able to adjust our evaluation for differences in in varying 409 submissions, we are unable to assess the underlying impact of modeling approaches on performance since we 410 do not have the granular details on forecast methods and how they evolved over time for all team forecasts. For 411 example, the LNQ-ens1, which outperformed all other forecasts by most metrics, only submitted forecasts for 412 approximately two thirds of the analysis period and stopped in June 2021 (prior to the Delta wave). The model is 413 described as a combination of three machine learning models, leveraging other embedded models and datasets, 414 with weights that "are chosen by hand each week based on performance in the previous week" (see LNQ-ens1 415 metadata, https://github.com/reichlab/covid19-forecast-

416 hub/blob/b12f916abc859bf59ea584b64f53afc2982042fd/data-processed/LNQ-ens1/metadata-LNQ-ens1.txt, at

417 (44)). The ensemble approach used in the LNQ-ens1 model building likely contributed to the overall 418 performance. However, several other ensemble models had lower performance than the LNQ-ens1 model; we 419 are unable to assess whether LNQ-ens1 performance gains were due to a particular component model or 420 dataset, the hand weighting procedure, or something else. The brief descriptions submitted to the COVID-19 421 Forecast Hub, such as for the LNQ-ens1, must include a summary of the methods used and may indicate a 422 variety of unique features such as input data, parameters, model fitting, etc. (44). However, the level of detail 423 provided in these descriptions varies between teams, and we do not have enough information to determine 424 which aspects of individual models were important determinants of forecast performance. To elucidate 425 associations between modeling approaches and forecast skill, additional research is needed. Future work to 426 support improved forecasting will require assessing the impact of specific features (e.g., through ablation 427 analyses) using retrospective, sable data systems and retrospective evaluation of the full forecasting process 428 (e.g., from data wrangling to final forecast production).

429

430 Infectious disease forecasting continues to present many challenges and opportunities for improving outbreak 431 response. Forecasts should be leading indicators of future activity and, while the COVID-19 case ensemble 432 forecasts were good leading indicators at many points in time; they were unreliable, especially during periods of 433 rapid change. Case data were integrated in COVID-19 mortality forecasts, which proved to be more reliable, 434 likely in part due to reported cases being leading indicators of reported deaths (18,45). However, because 435 deaths are a lagging indicator, death forecasts are less useful for short-term outbreak responses. Evaluation of 436 the case forecasts provided insight on limitations of early forecasts and research avenues for improving them. 437 These insights and the real-time forecasts provided by this effort were the product of large-scale collaboration 438 between researchers and public health responders to confront the COVID-19 pandemic. Learning from and 439 improving forecasting for COVID-19, other infectious diseases, and future pandemics will benefit from 440 continuing and expanding these collaborative efforts.

441

442 Methods

443 The US COVID-19 Forecast Hub (46) is a consortium of researchers that develop and share forecasts of COVID-19 444 reported cases, hospitalizations, and deaths with the goal of leveraging information from individual models that 445 predict the near-term burden of COVID-19 in the United States. Teams that submitted models to the US COVID-446 19 Forecast Hub used a wide variety of methodology and data (S1, Table S1). Beyond serving as a repository for 447 forecasts, submitted data were also aggregated by scientists at the COVID-19 Forecast Hub to generate two 448 models that we included in this analysis: the COVIDhub-4_week_ensemble and the COVIDhub-449 trained ensemble. Since the beginning of the COVID-19 Forecast Hub, the quantile predictions from each week's 450 submitted models were used as input data for the COVIDhub-4 week ensemble. Ensemble aggregation 451 methods evolved over time; for this analysis period, the ensemble forecast was calculated as the median across 452 forecasts from all models at each quantile level. Additionally, beginning on February 1, 2021, the COVID-19 453 Forecast Hub also generated a weighted ensemble (COVIDhub-trained ensemble). Models were selected for 454 weighted ensemble inclusion based on their past performance over various window period and given a weight 455 prior to aggregation. The methodology evolved over time and details are available on the model's metadata file 456 on the COVID-19 Forecast Hub GitHub repository (see Data and code availability and reporting guidelines).

457

The COVID-19 Forecast Hub, and death forecasts submitted to the Hub have been described in detail elsewhere (8,17,18). The Hub's incident COVID-19 case forecasts, which were first solicited in July 2020, have similar submission requirements to the death forecasts. Important differences include an expanded geographical scale (national; state, territory, and DC; and county levels) and reduced number of required quantiles in the probability distribution (7 quantiles in total: 0.025, 0.10, 0.25, 0.50, 0.75, 0.90, and 0.975). Predictions for weekly incident COVID-19 cases can be submitted for up to 8 weeks in the future, although our analysis only includes predictions made for 1-4 weeks into the future.

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466	We evaluated submitted forecasts between July 28, 2020, and December 21, 2021 (2020 epi week [EW] 31 –
467	2021 EW 51), which encompasses 73 weeks. Because forecasts were submitted at multiple geographic scales,
468	we conducted separate analyses for 1) national forecasts, 2) state, territory, and DC forecasts, 3) county level
469	forecasts, and 4) sets of team forecasts for all three geographic scales. When appropriate, we compared forecast
470	performance to that of a naïve model, created by the COVID-19 Forecast Hub, the COVIDhub-baseline. The
471	COVIDhub-baseline model, created each week, was designed to be a neutral model to provide a simple
472	reference point of comparison for all models. This baseline model forecasts a predictive median incidence equal
473	to the number of reported cases in the most recent week, with uncertainty based on the empirical distribution
474	of previous differences between the median and observed values (18).
475	
476	Inclusion criteria
477	Teams were included in the evaluation when they submitted forecasts with a complete set of quantiles for each
478	1- through 4-week ahead target predictions. Additionally, teams must have met the following inclusion criteria:
479	1. had predictions for at least 50 locations (states, territories, or DC) for the state, territory, and DC level
480	analyses; and for at least 75% of counties included in each population size quantile per submission week
481	for the county-level analyses;
482	2. had submissions for at least 50% of the weeks included in the analysis period per location forecasted.
483	
484	Teams meeting these inclusion criteria, and their submissions over time, are depicted in S1, Figure S1.
485	
486	Ground Truth
487	Forecasts were evaluated against the reported COVID-19 case reports collated by the Johns Hopkins Center for

488 Systems Science and Engineering (CSSE) (47). To calculate weekly incident reported cases, we subtracted the

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cumulative count for each Saturday from the cumulative count for the next Saturday, such that each incident weekly count reflects the number of cases reported from Sunday through Saturday in a given week. We aggregated reported counts from smaller geographic units into their larger unit. For example, counts in a given state are the aggregate of the county level reported counts and national counts are the sum of all states, territories, and DC.

494

495 CSSE reports data in real-time. Thus, data may be revised if the reporting health system makes public updates to 496 their surveillance data. At times, such revisions may result in negative daily counts or in increases to case counts 497 if the date of cases is shifted from one day to another or the definition of a reportable case is changed. We 498 examined the percent change between the first reported cases in each state, DC, and territory per date relative 499 to the counts in the surveillance file from April 2, 2022. We also assessed the influence of revised data on the 500 final model outcomes (see S2) and the presence of negative case counts in the timeseries. Less than 1 percent of 501 time points in the analysis period had negative daily case counts in the largest US counties. Negative counts 502 were observed at the state/territory level only twice: in Missouri during the week of April 17, 2021, and Virgin 503 Islands during the week ending October 10, 2020. The state of Florida reported 0 cases on November 27, 2021. 504 We excluded all weeks and locations with negative counts as well as the week with 0 incidence in Florida in our 505 primary analyses.

506

Additionally, we also examined whether a reported case count was an outlier in the case trend for each state. Anomalies in case data trends have not been uncommon throughout the pandemic, as reporting entities have uploaded large batches of surveillance data on a single day. To assess whether cases were outside of the expected range of reported cases over time, we applied three outlier detection algorithms, each with a 21-day window: a rolling median, a seasonal trend decomposition, and a seasonal trend decomposition without a seasonality term. We then classified a given count as an outlier if it was detected as such by at least two of the three algorithms. Using these data, we ran several sensitivity analyses to assess the likely impact of anomalous

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data points on model performance. Sensitivity analyses examining the robustness of our findings to reporting

- 515 anomalies are presented in S2.
- 516

517 Additional information about the CSSE data, and revisions to the dataset, is publicly available on a GitHub

- 518 repository:
- 519 https://github.com/CSSEGISandData/COVID-19/tree/master/csse_covid_19_data.
- 520

521 Forecast locations

522 Forecasts for incident cases were submitted for the national level, 50 states, 5 territories (American Samoa, 523 Guam, the Northern Mariana Islands, Puerto Rico, and the US Virgin Islands), the DC, and 3,142 US counties. We 524 excluded two counties in Alaska because they were not forecasted by most sets of team forecasts (Federal 525 Information Processing Standard code 02063 and 02066 were excluded). Because fewer teams submitted 526 forecasts for American Samoa, Guam, the Northern Mariana Islands, we excluded these territories from the 527 analysis. Some teams treated DC as both a county and a jurisdiction, so we excluded DC from the county 528 forecasts. In addition, because county population size and transmission are correlated and case counts and 529 forecast performance are also correlated, we grouped counties into 5 quantiles based on their population sizes, 530 with cut points at 8,908; 18,662; 36,742; and 93,230 people; most analyses used forecasts from the quantile 531 with the largest population size (n=628). We hypothesized that small counties would be more likely to have 532 better forecast accuracy because they had zero or very few reported cases. We thus chose to stratify counties by size to minimize any bias from aggregation. Performance results for most county forecasts are presented in S3. 533

534

535 **Defining epidemic phases**

For every state and DC, we independently classified each forecast week based on the estimated time-varying reproduction number (R_t) for that given week. State-level R_t estimates were obtained from

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538 <u>https://github.com/epiforecasts/epiforecasts.github.io</u> (48). We extracted the R_t estimate for the Wednesday of 539 each week from all available files. Because R_t estimates were updated on a rolling basis in near real time, there 540 were multiple estimates generated for the same date; we calculated the median estimated R_t per date for the 541 upper and lower 90% credible interval and the median value (August 1, 2020 – January 15, 2022, or 2020 EW 31 542 – 2022 EW 2, reflecting 77 weeks in total). Each forecast week was then classified into one of the following 543 categories based on the R_t estimates: *increasing, peak, decreasing, nadir.*

544

545 Increasing and decreasing phases reflect weeks in which Rt had a 90% probability of being greater than or less 546 than 1.0, respectively. There were several periods of rapid transmission in certain jurisdictions where R_t dipped 547 above/below the 1.0 threshold but did not remain on an upward or downward trajectory. Thus, we classified 548 weeks between two increasing phases as *increasing* and weeks between two decreasing periods as *decreasing*. 549 Weeks between increasing and decreasing phases were classified as *peaks*, whereas *nadirs* were defined as 550 periods between decreasing and increasing phases. Periods at the beginning or the end of an analysis period 551 were classified as a continuation of whichever phase preceded or followed them. Graphical depictions of Rt are 552 provided in S4 and show general concordance between R_t and reported cases.

553

554 Evaluation methodology

555 We evaluated probabilistic forecast accuracy using two different metrics, empirical prediction interval coverage 556 rates and weighted interval scores (WIS) (49). Coverage was calculated by determining the frequency with which 557 the prediction interval contained the eventually observed outcome for the 50%, 80% and 95% intervals. WIS 558 reflects a weighted estimate of sharpness (i.e., the range of the predicted interval) and calibration (i.e., precision 559 or error) across the three prediction intervals and the median prediction, with higher WIS and indicating lower 560 forecast skill. Importantly, WIS is highly correlated with the magnitude of observed and forecasted values. We 561 used mean absolute WIS to assess forecast accuracy over time and mean relative WIS (rWIS) to access forecast accuracy over space. Relative WIS was estimated by calculating the geometric mean of WIS across all sets of 562

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team forecasts and scaling that value to the WIS of a naïve model, the COVID-hub baseline. This approach eases interpretation, where values greater than 1.0 reflected worse accuracy than the baseline model and values below 1.0 reflected better model performance. Additionally, the pairwise relative comparison helps account for missing forecasts. Both coverage and WIS have been described in detail elsewhere (18,49). Horizon specific results for national, state/territory/DC, and large counties are presented in S5.

568

To assess the association between WIS and epidemic phase for each team, we fitted separate Gaussian 569 570 generalized estimating equation (GEE) models per team (equation 1) with an independent working correlation 571 structure at the state-level. This structure assumes that observations are not correlated over time in a state 572 (denoted as l in the equations below). Cases and weighted interval scores were log transformed and then 573 standardized (subtracting the mean and dividing by the standard deviation) prior to fitting the model, as this 574 transformation yielded more computationally and numerically stable estimates. We define those resulting 575 variables as stdWIS and stdCases. The expected value for a standardized WIS for time (t) and location (l), with 576 forecasts from a given team's model, is as follows:

577

578
$$\log\left(stdWIS_{t,l,h}\right) = \beta_0 + \alpha_{p[t,l]} + \gamma_h + ns\left(\log\left(stdCases_{t,l,h}\right)\right) + \epsilon_{t,l,h}$$
(1)

579

580 Where p[t, l] is an index that reflects the phase of each time (t) and location (l), (h) is the horizon of the 581 forecast in weeks, and $ns(\cdot)$ represents a natural spline with two degrees of freedom. Using a regression model 582 allows us to summarize patterns of overall average performance between teams while accounting for high 583 correlation and variation in the scores. Comparisons of rWIS, in contrast, do not allow for formal inference on 584 the differences in performance between teams. Prior to applying this regression model structure, our model 585 building approach included exploratory analysis of several structures appropriate for longitudinal analysis. We 586 examined model residuals, influential observations, goodness of fit metrics, and the impact of changing the 587 functional form of the variables included in the model.

588

589 The inclusion of reported cases in models permitted flexible adjustment for the wide range in cases between 590 and within jurisdictions, which led to a wide range of possible WIS values, as WIS values tend to be higher when 591 counts are higher. Expected WIS values were computed by first obtaining a marginal mean from the GEE model 592 and then undoing the transformations by exponentiating and un-standardizing the marginal mean. This was 593 done separately for each team for all phases and for each team and each phase individually (see S6 for 594 estimated team-specific marginal mean WIS relative to reported case counts). Additionally, we calculated 595 whether the 80% confidence interval (based on Gaussian distributional assumptions) for each team's expected 596 WIS outcome (on the log-scale and normalized, as described above) was less than the baseline model for all 597 phases (i.e., the marginal mean WIS for the baseline model).

598

599 Data and code availability and reporting guidelines

The forecasts from models used in this paper are available from the COVID-19 Forecast Hub GitHub repository (<u>https://github.com/reichlab/covid19-forecast-hub</u>) (8) and the Zoltar forecast archive (<u>https://zoltardata.com/project/44</u>) (50). The code used to generate all figures and tables in the manuscript is available in a public repository

604 (https://github.com/cdcepi/Evaluation-of-case-forecasts-submitted-to-COVID19-Forecast-Hub). All analyses 605 were conducted using the R language for statistical computing (v 4.0.3) (51), and the following packages were 606 used for the main analyses: *scoringutils* (52), *covidhubUtils* (53), *geepack* (54). Additionally, we included the 607 EPIFORGE 2020 reporting guideline checklist in S7 to indicate each page in this evaluation that corresponds to 608 each specific recommendation (15).

609

610 This activity was reviewed by CDC and was conducted consistent with applicable federal law and CDC policy.

25

611 **CDC disclaimer**: The findings and conclusions in this report are those of the authors and do not necessarily 612 represent the official position of the Centers for Disease Control and Prevention.

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763 Supporting information captions

764 Supporting Information 1: Team submissions, methods, and data

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SI Figure 1.1. Forecasts submitted over time at the national, state-territory-DC level in panel A and at the country scale in Panel B. The number of forecasted locations submitted each week nationally or at the state, territory and DC level is included, while the country level forecast submissions shows the percent of counties per quantile that were submitted each week. Sets of team forecasts meeting the inclusion criteria for this main analysis are labeled with an asterisk (*).

771

S1 Table 1.1. List of models evaluated, including sources for case, hospitalization, death, demographic, and mobility data when used as inputs for the given model. We evaluated 26 models contributed by 24 teams. The COVIDhub team submitted three models including the baseline model and the ensemble model. A brief description is included for each model, with a reference where available. The last column indicates whether the model made assumptions about how and whether social distancing measures were assumed to change during the period for which forecasts were made.

778 Supporting Information 2: Revision and outlier sensitivity analyses

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S2 Figure 2.1. To assess the influence of data revisions on our evaluation of forecast skill, we compared daily
 differences in cumulative reported cases during the week they were first reported to reported case counts for

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the same week in the complete data as of April 2, 2022. In total 721 weeks had at least one day with a revised case count (17% of all weeks, n=4,241 weeks) and revisions occurred in 43 of 51 jurisdictions. These jurisdiction specific plots compare cases reported as of the date in the subtitle (in red) compared to cases reported as of April 2, 2022 (in black).

786

787 S2 Figure 2.2. After identifying weeks with revised case counts, we then excluded them from the dataset and 788 reran the GEE models and estimated the marginal mean Weighted Interval Score (WIS). Panel A shows the 789 estimated marginal mean WIS and 95% confidence intervals for mean cases from team-specific GEE models 790 for all 48 jurisdictions from this sensitivity analysis. The 95% confidence intervals for the COVIDhub-baseline 791 model are shown in dashed red vertical lines. Panel B presents each team's estimated marginal mean WIS per 792 phase, scaled to the COVIDhub-baseline model's estimated marginal mean WIS for all epidemic phases, using 793 the dataset with excluded week. Teams with higher estimated marginal mean WIS values (i.e., greater than 794 1.0) are presented in shades of orange while teams with lower estimated marginal mean WIS (i.e., less than 795 1.0) are shown in shades of green. Team forecasts are denoted with an asterisk (*) if the 80% confidence 796 interval of the expected WIS outcome (normalized and on the log scale) was estimated by a model to be lower 797 than the expected WIS of the COVIDhub-baseline model for all phases.

798

52 Figure 2.3. Outliers were defined as non-revised reported case counts that were outside of the expected range by at least two of the three algorithms: a rolling median, a seasonal trend decomposition, and a seasonal trend decomposition without a seasonality term. Each method used a 21-day window. Approximately three percent of weeks (686 of 27,489 total weeks in the analysis period) had at least one day of reported cases identified as an outlier.

804

805	³³ Supporting Information 3: Incident COVID-19 case forecasts were submitted for all US counties. The
806	plots shown here depicted average, scaled pairwise Weighted Interval Score (WIS; see <i>Methods</i> for
807	description), 95% coverage, and submissions (S3 Figure 3.1), average 50%, 80% and 95% coverage for
808	eligible submitted forecasts (S3 Figure 3.2), and average WIS and 95% coverage over time (S3 Figure
809	3.2). Each figure shows spatial disaggregated results, with increasing population size and quantile
810	numbers. For example, counties with the smallest population are grouped in Quantile 1 and the
811	largest population sizes are grouped in Quantile 5. The following teams are included in these figures:
812	CEID-Walk, LNQ-ens1, Microsoft_DeepSTIA, COVIDhub-4_week_ensemble, COVIDhub-
813	trained_ensemble, COVIDhub-baseline, CU-select, FAIR-NRAR, FRBSF_Wilson-Econometric,
814	lowasStateLW-STEM, JHU_IDD-CovidSP, JHU_CSSE-DECOM, JHUAPL-Bucky, LANL-GrowthRate, LNQ-
815	esn1, UVA-Ensemble.
816	
017	S2 Figure 2.1. Demont of weaks with complete submissions for all sets of team foresets, scaled, painwise
017	55 Figure 5.1. Percent of weeks with complete submissions for an sets of team forecasts, scaled, pairwise
818	relative Weighted Interval Score (rWIS), 95% coverage, and by geographical scale of submitted forecasts.
817 818 819	relative Weighted Interval Score (rWIS), 95% coverage, and by geographical scale of submitted forecasts. Teams are sorted by increasing rWIS values.
817 818 819 820	relative Weighted Interval Score (rWIS), 95% coverage, and by geographical scale of submitted forecasts. Teams are sorted by increasing rWIS values.
817 818 819 820 821	relative Weighted Interval Score (rWIS), 95% coverage, and by geographical scale of submitted forecasts. Teams are sorted by increasing rWIS values. S3 Figure 3.2. Expected and observed coverage rates aggregated over time and horizon for county forecasts.
817 818 819 820 821 822	relative Weighted Interval Score (rWIS), 95% coverage, and by geographical scale of submitted forecasts. Teams are sorted by increasing rWIS values. S3 Figure 3.2. Expected and observed coverage rates aggregated over time and horizon for county forecasts. The dashed line represents optimal expected-coverage. Team forecasts that outperformed the COVIDhub-
 817 818 819 820 821 822 823 	 relative Weighted Interval Score (rWIS), 95% coverage, and by geographical scale of submitted forecasts. Teams are sorted by increasing rWIS values. S3 Figure 3.2. Expected and observed coverage rates aggregated over time and horizon for county forecasts. The dashed line represents optimal expected-coverage. Team forecasts that outperformed the COVIDhub-4_week_ensemble model at all coverage levels are labeled on the right hand side of the plots.
 817 818 819 820 821 822 823 824 	 relative Weighted Interval Score (rWIS), 95% coverage, and by geographical scale of submitted forecasts. Teams are sorted by increasing rWIS values. S3 Figure 3.2. Expected and observed coverage rates aggregated over time and horizon for county forecasts. The dashed line represents optimal expected-coverage. Team forecasts that outperformed the COVIDhub-4_week_ensemble model at all coverage levels are labeled on the right hand side of the plots.
 817 818 819 820 821 822 823 824 825 	 S3 Figure 3.1. Percent of weeks with complete submissions for all sets of teach forecasts, scaled, pairwise relative Weighted Interval Score (rWIS), 95% coverage, and by geographical scale of submitted forecasts. Teams are sorted by increasing rWIS values. S3 Figure 3.2. Expected and observed coverage rates aggregated over time and horizon for county forecasts. The dashed line represents optimal expected-coverage. Team forecasts that outperformed the COVIDhub-4_week_ensemble model at all coverage levels are labeled on the right hand side of the plots. S3 Figure 3.3. Mean Weighted Interval Score (WIS) over time, aggregated by geographic units and forecast

827 dashed vertical line in all panels shows the date that public communication of the case forecasts was paused.

828 The black, dashed horizontal line in panels B show nominal 95% interval coverage

829

830	Supporting Information 4: Estimated time-varying reproduction number and epidemic phase
831	classifications. For each state, the top panel shows the median R_t and median upper and lower 90%
832	credible interval over time in red. The bottom panel shows reported case counts over time. Both
833	plots have vertical bands representing the epidemic phase of each forecast week: <i>increasing, peak</i> ,
834	decreasing, nadir.
835	
0.2.6	
836	Supporting information 5: Each location specific forecast submitted to the COVID19 Forecast Hub
837	included at least 4 weeks of future predictions. Here, we present disaggregated 1 and 4 week ahead
838	predictions of model performance for each team model that submitted national and
839	state/territory/DC forecasts and were included in the main analyses. Specific plots include the
840	average 50%, 80% and 95% coverage for eligible submitted forecasts (S5 Figure 5.1), average
841	absolute Weighted Interval Score (WIS) and 95% coverage over time (S5 Figure 5.2), and scaled,
842	pairwise rWIS by location (S5 Figure 5.3)
843	
844	S5 Figure 5.1. Expected and observed coverage rates aggregated for 1 and 4 week ahead forecasts over time
845	for national forecasts in A, state/territory/DC forecasts in B, the largest country forecasts in C. The dashed line
846	represents optimal expected-coverage. Teams that outperformed the COVIDhub-4_week_ensemble model at
847	all coverage levels are labeled on the right-hand side of the plots.

848

S5 Figure 5.2. Mean Weighted Interval Score (WIS) over time for 1 and 4 week ahead forecasts, aggregated by
 geographic units, and 95% coverage over time for 1 and 4 week ahead forecasts, aggregated by geographic

35

851	units. The black, dashed vertical line in all panels shows the date that public communication of the case
852	forecasts was paused. The black, dashed horizontal line in panels D, E, and F show nominal 95% interval
853	coverage. Teams that submitted national forecasts are presented in A. and D., state/territory/DC forecasts
854	presented in B. and E., and teams that submitted large county level forecasts are presented in C. and F.
855	
856	
857	S5 Figure 5.3. Scaled, pairwise relative Weighted Interval Score (rWIS; see <i>Methods</i> for description) for all
858	teams that submitted national and state/territory/DC forecasts by location for 1 and 4 week ahead horizon.
859	National estimates are displayed first, followed by jurisdictions in alphabetical order. Teams are displayed by
860	decreasing average rWIS across all forecast horizons and locations.
861	
862	Supporting Information 6: Each team model's estimated marginal mean Weighted Interval Score
863	(WIS) over range of reported case counts per epidemic phase. Marginal mean WIS was estimated
864	from GEE model results and reflect values across the 95% confidence interval of mean reported
865	cases. Case counts differ per team model as each team forecasted a different set of locations over a
866	different range of possible dates.
867	

Supporting Information 7: EPIFORGE 2020 guidelines outline 19 recommended reporting items for
 epidemic forecasting and prediction research (15). These items are included in the checklist below,
 which also include the page number where each item is described or presented within this
 evaluation.

36

872 Main Text Figures

873

Figure 1. Weekly incident reported COVID-19 cases per 100K population, nationally (in black) and per state/territory/DC (in gray), over time in panel A. Panel B shows a subset of COVIDhub-4_week_ensemble forecasts (in green) over time, with the median predictions represented as lines and points and the 95% prediction intervals in bands. Reported incident cases (counts per week) are shown in gray. In both plots, the black, dashed vertical line shows the date that public communication of the case forecasts was paused.



Figure 2: Percent of weeks with complete submissions for all sets of team forecasts, scaled, pairwise relative Weighted Interval Score (rWIS; see Methods for 0

description), observed 95% prediction interval coverage, by geographical scale of submitted forecasts. Teams are sorted by increasing state/territory/DC rWIS 1

values. 2

A. Percent of Weeks Submitted				B. Relative Weighted Interval Score			C. 95% Coverage		
LNQ-ens1	0.62	0.62		0.65	0.59		0.97	0.96	
COVIDhub-4_week_ensemble	1.00	1.00		0.81	0.80	-	0.80	0.76	
USC-SI_kJalpha	1.00	1.00		0.83	0.68	-	0.60	0.78	
LANL-GrowthRate	0.84	0.84		0.86	0.72		0.85	0.85	
Microsoft-DeepSTIA	0.69	0.69		0.96	0.95	-	0.64	0.56	
COVIDhub-trained_ensemble	0.53	0.53		0.97	1.08	-	0.80	0.73	
CU-select-	0.96	0.96		0.99	0.92		0.67	0.77	
COVIDhub-baseline	1.00	1.00		1.00	1.00	-	0.77	0.72	
BPagano-RtDriven	0.84	0.84		1.00	0.83	-	0.72	0.71	
JHUAPL-Bucky	0.93	0.93		1.01	0.75		0.64	0.79	
IEM_MED-CovidProject	0.74	0.74		1.03	0.80	-	0.80	0.89	
MOBS-GLEAM_COVID	0.57	0.57		1.05	0.87	-	0.73	0.77	
JHU_CSSE-DECOM-	0.82	0.70		1.06	0.92		0.57	0.67	
CEID-Walk	0.84	0.84		1.08	1.10	-	0.64	0.52	
Karlen-pypm	0.97	0.97		1.10	1.34	-	0.77	0.43	
UVA-Ensemble	0.77	0.77		1.11	1.11	-	0.84	0.91	
JHU_UNC_GAS-StatMechPool	0.73			1.13		-	0.86		
JHU_IDD-CovidSP	0.82	0.82		1.14	1.40	-	0.88	0.43	
RobertWalraven-ESG	1.00	0.99		1.21	1.36	-	0.46	0.40	
lowaStateLW-STEM-	0.53	0.53		1.32	1.19	-	0.20	0.23	
Covid19Sim-Simulator	0.72	0.72		1.34	1.23	-	0.25	0.26	
UCLA-SuEIR	0.64	0.64		1.39	1.12	-	0.44	0.47	
CovidAnalytics-DELPHI	1.00	1.00		1.45	1.24	-	0.63	0.57	
L	State/Territory/DC Forecasts	National Forecasts		State/Territory/DC Forecasts	National Forecasts	_	State/Territory/DC Forecasts	National Forecasts	
	% Weeks o.e	0.7 0.8 0.9 1.0	Rela	tive Weighted Interval	Score 0.6 0.8 1.0 1.2	1	95% Coverage 0.2	0.4 0.6 0.8	

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- 38
- Figure 3: Expected and observed coverage rates for central 50%, 80% and 95% prediction intervals aggregated over time
- 35 and horizon for national forecasts (panel A), state/territory/DC forecasts (panel B), the largest county forecasts (panel
- 36 C). The dashed line represents optimal expected-coverage. Team forecasts that had closer to nominal coverage than the
- 37 COVIDhub-4_week_ensemble model at all three coverage levels are labeled on the right side of the plots.



38

Figure 4: Scaled, pairwise relative Weighted Interval Score (rWIS) (see *Methods* for description) by spatial scale for sets of team forecasts that submitted forecasts for the US nation, states/territories/DC, and all US counties. WIS is averaged across all horizons. The COVIDhub-baseline model has, by definition, a rWIS of 1 (horizontal dashed line). Teams are ordered by increasing state/territory/DC rWIS with the most accurate model on the left. Points for each team are staggered horizontally to show overlapping WIS values.

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Figure 5: Scaled, pairwise relative Weighted Interval Score (rWIS; see *Methods* for description) by location for national

38 and state/territory/DC forecasts, averaged across all horizons through the entire analysis period. National estimates are

- 39 displayed first, followed by jurisdictions in alphabetical order. Team forecasts are ordered by increasing average
- 00 state/territory/DC rWIS.



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)1

- Figure 6: Forecast accuracy over time, aggregated by geographic units, forecast horizon, and prediction date. Panels A-C show average Weighted Interval Score (WIS); panels D-F show 95% prediction interval coverage. The black, dashed vertical line in all panels shows the date that public communication of the case forecasts was paused. The black, dashed horizontal line in panels D-F shows nominal 95% interval coverage. National level forecasts are presented in A and D,
- 37 state/territory/DC forecasts in B and E and large county level forecasts in C and F.



Panels A, D, B and E include: LNQ-ens1, Microsoft_DeepSTIA, COVIDhub-4_week_ensemble, USC-SI_kJalpha, CU-select, LANL-GrowthRate, JHU_CSSE-DECOM, COVIDhub-trained_ensemble, COVIDhub-baseline, Karlen-pypm, BPagano-RtDriven, JHUAPL-Bucky, UVA-Ensemble, IEM_MED-CovidProject, CEID-Walk, Covid19Sim-Simulator, IowasStateLW-STEM, UCLA-SuEIR, JHU_IDD-CovidSP, RobertWalraven-ESG, MOBS-GLEAM_COVID, and CovidAnalytics-DELPHI.

Panels C and F include: LNQ-ens1, COVIDhub-4_week_ensemble, CU-select, LANL-GrowthRate, COVIDhub-trained_ensemble, COVIDhub-baseline, JHUAPL-Bucky, UVA-Ensemble, CEID-Walk, JHU_UNC_GAS-StatMechPoole, IowasStateLW-STEM, JHU_IDD-CovidSP, UMass-MechBayes, FAIR-NRAR, FRBSF_Wilson-Econometric.

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Figure 7. Estimated marginal mean Weighted Interval Score (WIS) and 95% confidence intervals for mean cases from 11 12 team-specific GEE models for all 51 jurisdictions (Panel A). The 95% confidence intervals for the COVIDhub-baseline model are shown in dashed red vertical lines. Panel B presents each team's estimated marginal mean WIS per phase, 13 14 scaled to the COVIDhub-baseline model's estimated marginal mean WIS for all epidemic phases. Teams with higher 15 estimated marginal mean WIS values (i.e., greater than 1.0) are presented in shades of orange while teams with lower 16 estimated marginal mean WIS (i.e., less than 1.0) are shown in shades of green. Forecasts for a team in a particular 17 phase are marked with an asterisk (*) if the 80% confidence interval of the expected WIS outcome (normalized and on the log scale) was estimated by a model to be lower than the expected WIS of the COVIDhub-baseline model for all 18 19 phases.



*80% CI is lower than baseline

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